## **REGRESSION MODELS**

*One approach:* Use theoretical considerations to develop a model for the mean function or other aspects of the conditional distribution.

The next two approaches require some terminology:

- Error: e|x = Y|(X = x) E(Y|X = x)= Y|x - E(Y|x) for short
  - So Y|x = E(Y|x) + e|x (Picture this ...)
  - elx is a random variable
  - E(e|x) = E(Y|x) E(Y|x)) = E(Y|x) E(Y|x) = 0
  - Var(e|x) =
  - The distribution of elx is

## Second approach:

Bivariate Normal Model: Suppose X and Y have a bivariate normal distribution.

### Recall:

Ylx is normal
E(Y| x) = μ<sub>Y</sub> + ρ σ<sub>Y</sub>/σ<sub>X</sub> (x - μ<sub>X</sub>) (linear mean function)
Var(Ylx) = σ<sub>Y</sub><sup>2</sup>(1 - ρ<sup>2</sup>) (constant variance)

Thus:

• E(Y|x) = a + bx•  $Var(Y|x) = \sigma^2$ where b =

a =

 $\sigma^2 =$ 

Implications for elx:

•  $e|x \sim$ 

## "The" Simple Linear Regression Model

#### Version 1:

Only one assumption: E(Y|x) is a linear function of x.

Typical notation: $E(Y|x) = \eta_0 + \eta_1 x$ (or  $E(Y|x) = \beta_0 + \beta_1 x$ )Equivalent formulation: $Y|x = \eta_0 + \eta_1 x + e|x$ Interpretations of parameters:(Picture!) $\eta_1$ :(if  $\eta_1$ )

 $\eta_0$  :

(if ...)

## When model fits:

- X, Y bivariate normal
- Other situations Example: Blood lactic acid Why is this not bivariate normal?
- Model might also be used when mean function is not linear, but linear approximation is reasonable.

### **Version 2**: *Two assumptions*:

- 1.  $E(Y|x) = \eta_0 + \eta_1 x$  (linear mean function)
- 2.  $Var(Y|x) = \sigma^2$  (constant variance)

Equivalent formulation: 1'.  $E(Y|x) = \eta_0 + \eta_1 x$  (linear mean function) 2':  $Var(e|x) = \sigma^2$  (constant error variance) raw a picture!

[Draw a picture!]

Situations where the model fits:

- If X and Y have a bivariate normal distribution.
- Credible (at least approximately) in many other situations as well, for transformed variables if not for the original predictor. (i.e., it's often useful)

Until/unless otherwise stated, we will henceforth assume the Version 2 model -- i.e., we will assume conditions (1) and (2) (equivalently, (1') and (2').)

Thus we have *three parameters*:

 $\eta_0,\eta_1$  (which determine E(Ylx) and  $\sigma^2$  (which determines Var(Ylx)

**The goal**: To estimate  $\eta_0$  and  $\eta_1$  (and later  $\sigma^2$ ) from data.

*Notation*: The estimates of  $\eta_0$  and  $\eta_1$  will be called  $\hat{\eta}_0$  and  $\hat{\eta}_1$ , respectively. From  $\hat{\eta}_0$  and  $\hat{\eta}_{l}$ , we obtain an estimate

$$\hat{\mathbf{E}}(\mathbf{Y}|\mathbf{x}) = \hat{\eta}_0 + \hat{\eta}_1 \mathbf{x}$$

of E(Y|x).

*Note*:  $\hat{E}(Y|x)$  is the same notation we used earlier for the lowess estimate of E(Y|x). Be sure to keep the two estimates straight.

*More terminology*:

- We label our data  $(x_1, y_1), (x_2, y_2), \dots, (x_n, y_n)$ .
- $\hat{y}_i = \hat{\eta}_0 + \hat{\eta}_1 x_i$  is our resulting estimate  $\hat{E}(Y|x_i)$  of  $E(Y|x_i)$ . It is called the  $i^{th}$ *fitted value* or *i<sup>th</sup> fit. ê<sub>i</sub>* = y<sub>i</sub> - ŷ<sub>i</sub> is called the *i<sup>th</sup> residual.*

*Note*:  $\hat{e}_i$  (the residual) is analogous to but not the same as  $elx_i$  (the error). Indeed,  $\hat{e}_i$  can be considered an estimate of the error  $e_i = y_i - E(Y|x_i)$ .

Draw a picture:

**Least Squares Regression:** A method of obtaining estimates  $\hat{\eta}_0$  and  $\hat{\eta}_1$  for  $\eta_0$  and  $\eta_1$ 

Consider lines  $y = h_0 + h_1 x$ . We want the one that is "closest" to the data points  $(x_1, y_1)$ ,  $(x_2, y_2), \ldots, (x_n, y_n)$  collectively.

What does "closest" mean?

Various possibilities:

# 1. The usual math meaning: shortest perpendicular distance to point.

## Problems:

- Gets unwieldy quickly.
- We're really interested in getting close to y for a given x -- which suggests:
- 2. Minimize  $\sum d_i$ , where  $d_i = y_i (h_0 + h_1 x_i) =$  vertical distance from point to candidate line. (Note: If the candidate line is the desired best fit then  $d_i = \dots$ ) Problem: Some  $d_i$ 's will be positive, some negative, so will cancel out in the sum. This suggests:
- 3. Minimize  $\sum |d_i|$ . This is feasible with modern computers, and is sometimes done. Problems:
  - This can be computationally difficult and lengthy.
  - The solution might not be unique. Example:
  - The method does not lend itself to inference about the fit.

4. Minimize  $\sum d_i^2$ 

This works! See demo.

Terminology:

- $\sum_{i=1}^{\infty} d_i^2$  is called the *residual sum of squares* (denoted *RSS*( $h_0, h_1$ )) or the *objective function*.
- The values of  $h_0$  and  $h_1$  that minimize RSS( $h_0$ ,  $h_1$ ) are denoted  $\hat{\eta}_0$  and  $\hat{\eta}_1$ , respectively, and called the *ordinary least squares* (or *OLS*) *estimates* of  $\eta_0$  and  $\eta_1$