

WEIGHTED LEAST SQUARES

Model assumptions:

1. $E(Y|\underline{x}_i) = \underline{\eta}^T \underline{u}_i$ (linear mean function -- as for ordinary least squares)
2. $\text{Var}(Y|\underline{x}_i) = \sigma^2/w_i$, where the w_i 's are known, positive constants (called *weights*)
(Different from OLS!)

Observe:

- w_i is inversely proportional to $\text{Var}(Y|\underline{x}_i)$. This is sometimes helpful in getting suitable w_i 's.
- the w_i 's aren't unique – we could multiply all of them by a constant c , and divide σ by \sqrt{c} to get an equivalent model.

For WLS, the *error* is defined as

$$e_i = \sqrt{w_i} [Y|\underline{x}_i - \underline{\eta}^T \underline{u}_i] \quad (\text{Different from OLS!})$$

Then (exercise)

$$\begin{aligned} E(e_i) &= 0 \\ \text{Var}(e_i) &= \sigma^2 \end{aligned}$$

Reformulating (1) in terms of errors:

$$1': Y|\underline{x}_i = \underline{\eta}^T \underline{u}_i + e_i/\sqrt{w_i}$$

Note: WLS is not a universal remedy for non-constant variance, since weights are needed. But it is useful in many types of situations, e.g.,

- A. If $Y|\underline{x}_i$ is the sum of m_i independent observations, each with variance σ^2 , then
 $\text{Var}(Y|\underline{x}_i) = \underline{\hspace{2cm}}$, so we could take $w_i = \underline{\hspace{2cm}}$.
- B. If $Y|\underline{x}_i$ is the average of m_i independent observations, each with variance σ^2 , then
 $\text{Var}(Y|\underline{x}_i) = \underline{\hspace{2cm}}$, so we could take $w_i = \underline{\hspace{2cm}}$.
- C. Sometimes visual or other evidence suggests a pattern of how $\text{Var}(Y|\underline{x}_i)$ depends on x_i . For example, if it looks like $\sqrt{\text{Var}(Y|x_i)}$ is a linear function of x_i [Sketch a picture of this!], then we can fit a line to the data points (x_i, s_i) , where s_i = sample mean of observations with x value x_i . If we get
 $\hat{s}_i = \hat{\gamma}_0 + \hat{\gamma}_1 x_i$, try $w_i = \underline{\hspace{2cm}}$.
- D. Sometimes theoretical considerations may suggest a choice of weights. (e.g., theoretical considerations might suggest that the conditional distributions are Poisson, which implies that their variances are equal to their means. This would suggest taking $w_i = \underline{\hspace{2cm}}$.)

E. Weighted least squares is also useful for other purposes – e.g., in calculating the lowest estimate, lines are fit so that points at the ends of the range count less than points at the middle of the range.

A WLS model may be fit by least squares: Find $\hat{\underline{h}}$ to minimize the “weighted residual sum of squares”

$$\text{RSS}(\underline{h}) = \sum w_i (y_i - \underline{h}^T \underline{u}_i)^2$$
$$\hat{\underline{h}}$$
 is called the “WLS estimate” of the coefficients.

Comments:

- a. If all $w_i = 1$, we get _____.
- b. The larger w_i is, the more the i^{th} observation “counts” (and the _____ er the variance at x_i – think of the geese example.)
- c. $\text{RSS}(\underline{h}) = \sum [\sqrt{w_i} y_i - \underline{h}^T (\sqrt{w_i} \underline{u}_i)]^2$, so we could get $\hat{\underline{h}}$ by using OLS to regress the $\sqrt{w_i} y_i$'s on the $\sqrt{w_i} \underline{u}_i$'s, *but*, we would need to fit *without* an intercept, since the first component of $\sqrt{w_i} \underline{u}_i$ is. However, most software will fit a WLS if given the weights.

Example: Coin data.